

## THE TEST OF THE DISCRIMINATION FOR TWO POISSON PROCESSES

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In this paper we give first the test of the discrimination for two Poisson processes. Next we study the asymptotic normality of the statistics determining the test and the convergence rate in this normal asymptotic process.

### I - A RESTRICTED TEST PROBLEM

Let  $T$  be a separable metric space,  $\mathfrak{B}$  be the  $\sigma$ -field of all Borel subsets of  $T$ . Standard notations and concepts in [1] (cf. Chapters II, III) will be used throughout the paper. Assume that a realization  $\mu$  of the Poisson process  $Q$  is observed and we have to choose one of two possibilities:  $Q = Q_{\rho_0}$  or  $Q = Q_{\rho_1}$ , where  $Q_{\rho_0}, Q_{\rho_1}$  are Poisson processes with densities  $\rho_0, \rho_1$  ( $\rho_0, \rho_1$  are Radon measures on  $T$ ). Thus we have the following test problem:

$$(H) \left\{ \begin{array}{l} \text{Test of the hypothesis } Q = Q_{\rho_0} \\ \text{with the alternative hypothesis } Q = Q_{\rho_1} \end{array} \right.$$

In general, this problem has no testable solution by a simple observation and we are led to the following restricted test problem:

$$(H^K) \left\{ \begin{array}{l} \text{Test of the hypothesis } Q^K = Q_{\rho_0}^K \\ \text{with the alternative hypothesis } Q^K = Q_{\rho_1}^K \end{array} \right.$$

where  $K$  is a compact subset of  $T$  and  $Q^K, Q_{\rho_0}^K, Q_{\rho_1}^K$  are images of  $Q, Q_{\rho_0}, Q_{\rho_1}$  under the mapping  $\mu \rightarrow \mu^K$  ( $\mu^K$  is the restriction of  $\mu$  on  $K$ ).

Let  $\rho$  be a Radon measure on  $T$  such that

$$\rho_i \ll \rho \quad (i = 0, 1)$$

Then for every compact subset  $K$  of  $T$

$$\rho_i^K \ll \rho^K \quad (i = 0, 1).$$

Therefore, from [1] (Theorem 1, Chapter VI), we have:

$$Q_{\rho_i}^K \ll Q_{\rho}^K \quad (i = 0, 1)$$

with

$$\frac{dQ_{\rho_i}^K}{dQ_{\rho}^K}(\mu) = \exp \{ -(\rho_i(K) - \rho(K)) + \mu(\ln f_i) \},$$

$$f_i(t) = \frac{d\rho_i}{d\rho}(t) \quad (i = 0, 1).$$

By the Neymann — Pearson's basis Lemma (cf, [3], p. 95-96), it follows that the most powerful test with significance level  $\alpha$  of the problem  $(H_K)$  have the critical domain of the form

$$S_K^\alpha \left\{ \mu : \mu^K \left( \ln \frac{f_1}{f_0} \right) > c(\alpha) \right\},$$

where  $c(\alpha)$  is defined so that

$$Q_{\rho_0}^K \left\{ S_K^\alpha \right\} = \alpha.$$

Hence, the most powerful test with significance level  $\alpha$  of the problem  $(H_K)$  is completely determined by the statistics

$$Z_K(\mu) = \mu^K \left( \ln \frac{f_1}{f_0} \right).$$

Now we investigate the distribution of  $Z_K$ .

Since

$$(1) \quad \{Z_K > -\infty\} \Leftrightarrow \text{supp } \mu \subset K \cap \{f_1 > 0\},$$

and

$$(2) \quad \{Z_K < +\infty\} \Leftrightarrow \text{supp } \mu \subset K \cap \{f_0 > 0\},$$

its follows that

$$Q_{\rho_0} \{Z_K = -\infty\} = 1 - \exp \{-\rho_0(K \cap \{f_1 = 0\})\}; \quad Q_{\rho_0} \{Z_K < +\infty\} = 1,$$

$$Q_{\rho_1} \{Z_K = +\infty\} = 1 - \exp \{-\rho_1(K \cap \{f_0 = 0\})\}; \quad Q_{\rho_1} \{Z_K > -\infty\} = 1.$$

Further, from (1) we infer that under the hypothesis  $Q^K = Q_{\rho_0}^K$  the characteristic function of  $Z_K$  with  $Z_K > -\infty$  is of the form

$$\varphi_K^0(u) = \exp \{ \rho_0 [I_{K \cap \{f_1 > 0\}} (\exp(iu \ln \frac{f_1}{f_0}) - 1)] \}.$$

Analogously, from (2) it follows that under the alternative hypothesis  $Q^K = Q_{\rho_1}^K$ , the characteristic function of  $Z_K$  with  $Z_K < +\infty$ , is of the form

$$\varphi_K^1(u) = \exp \{ \rho_1 [I_{K \cap \{f_0 > 0\}} (\exp(iu \ln \frac{f_1}{f_0}) - 1)] \}.$$

These results give us all informations to determine the critical point  $c(\alpha)$  and the power of the test of the problem  $(H_K)$ .

## II — THE ASYMPTOTIC NORMALITY

Let  $K_1, K_2, \dots$  denote an increasing sequence of compact subsets of  $T$  such that  $\lim_{n \rightarrow \infty} K_n = T$ . Put

$$Z_n(\mu) = Z_{K_n}(\mu), \quad n = 1, 2, \dots$$

Denote by  $L_i(X)$  the distribution of the statistics  $X$  with respect to the probability measure  $Q_{\rho_i}$  ( $i = 0, 1$ ).

**THEOREM 1:** *Assume that the following conditions are satisfied*

1)  $\rho_0 \ll \rho_1$

2)  $\int_T (f^{1/2}(t) - 1)^2 \rho_1(dt) = \infty$       where  $f(t) = \frac{d\rho_0}{d\rho_1}(t)$ .

3) There exists  $M > 0$  such that

$$(\rho_0 + \rho_1) \{x: | \ln f(x) | > M\} < \infty$$

Then as  $n \rightarrow \infty$

$$L_0 \left( \frac{Z_n - a_n}{b_n} \right) \rightarrow \mathcal{N}(0,1)$$

where

$$a_n = \int_{K_n^M} \ln \frac{1}{f(t)} \rho_0(dt), \quad b_n = \left( \int_{K_n^M} \ln \frac{4}{f(t)} \rho_0(dt) \right)^{1/2}$$

$$K_n^M \cap K^M, \quad K^M = \{x: | \ln f(x) | \leq M\}$$

Proof. Note that  $Z_n$  can be decomposed as follows

$$Z_n = Z_n^{(1)} + Z_n^{(2)}$$

where

$$Z_n^{(1)} = \int_{K_n^M} \ln \frac{1}{f(t)} \mu(t), \quad Z_n^{(2)} = Z_n - Z_n^{(1)}$$

We shall prove that

$$(3) \quad L_0 \left( \frac{Z_n^{(1)} - a_n}{b_n} \right) \rightarrow \mathcal{N}(0,1)$$

and

$$Q_{\rho_0} \left\{ \frac{Z_n^{(2)}}{b_n} \rightarrow 0 \right\} = 1.$$

Since for  $Q_{\rho_0}$  — almost all  $\mu$  there exists a finite limit

$$Z^{(2)} = \lim_{n \rightarrow \infty} Z_n^{(2)}$$

it suffices to prove (3) and

$$(4) \quad \lim_{n \rightarrow \infty} b_n = \infty$$

From the condition 2) of the Theorem we have

$$(5) \quad \infty = \int_{K^M} \left(1 - \frac{1}{f^{1/2}}\right)^2 \rho_0(dt) + \int_{K^M} \left(1 - \frac{1}{f^{1/2}}\right)^2 \rho_1(dt)$$

where  $\overline{K^M} = \{x: | \ln f(x) | > M\}$ .

We can write

$$\overline{K^M} = K_{>}^M + \overline{K_{<}^M}$$

where

$$\overline{K_{>}^M} = \left\{ x : \frac{1}{f(x)} > e^M \right\}, \quad \overline{K_{<}^M} = \overline{K^M} \setminus \overline{K_{>}^M}.$$

Thus we have the estimation

$$(6) \quad \left(1 - \frac{1}{f^{1/2}}\right)^2 < \frac{1}{f} \quad \text{on } \overline{K_{>}^M}$$

and

$$(7) \quad \left(1 - \frac{1}{f^{1/2}}\right)^2 < 1 \quad \text{on } \overline{K_{<}^M}.$$

From (5) + (7) and from the condition 3) of the Theorem it follows that

$$(8) \quad \int_{\overline{K^M}} \left(1 - \frac{1}{f^{1/2}}\right)^2 \rho_0(dt) = \infty$$

Further, we see that (8) is satisfied if and only if one of the following conditions holds:

i) 
$$\int_{\overline{K^M} \cap \{| \ln f | \leq c\}} \left(1 - \frac{1}{f^{1/2}}\right)^2 \rho_0(dt) = \infty \text{ for every } c > 0$$

(ii) There exists  $c > 0$  such that

$$\int_{\overline{K^M} \cap \{| \ln f | > c\}} \left(1 - \frac{1}{f^{1/2}}\right)^2 \rho_0(dt) = \infty$$

From (i) it follows there exists  $c > 0$  such that if  $| \ln f | < c$  then

$$(9) \quad \ln^2 \frac{1}{f} > \left(1 - \frac{1}{f^{1/2}}\right)^2$$

which together with (8) implies (4).

To prove (3) let us denote

$$T_n(t) = \frac{1}{b_n} \ln \frac{1}{f(t)}.$$

Hence, the logarithm  $\psi_n(u)$  of the characteristic function of

$$Z_n = \frac{Z_n - a_n}{b_n} \text{ is of the form}$$

$$\psi_n(u) = \int_{\overline{K_n^M}} \exp(iu T_n(t) - 1 - iu T_n(t)) \rho_0(dt).$$

Thus we have the estimation

$$(10) \quad \left| \psi_n(u) + \frac{u^2}{2} \right| \leq \int_{K_n^M} \left| \exp(iuT_n(t) - 1 - iuT_n(t) + \frac{u^2}{2} T_n(t) \right| \rho_0(dt) \leq \\ \leq \frac{M |u|^3}{b_n}$$

which together with (3) implies (3).

Theorem 1 is proved.

Analogously we can prove the following

**THEOREM 2.** Assume that following conditions are satisfied

1')  $\rho_1 \ll \rho_0$

2')  $\int_{\Gamma} (g^{v^2}(t) - 1)^2 \rho_0(dt) < \infty$  where  $g(t) = \frac{d\rho_1}{d\rho_0}(t)$ .

3) There exists  $M > 0$  such that

$$(\rho_0 + \rho_1) \{x : |\ln g(x)| > M\} < \infty$$

Then as  $n \rightarrow \infty$

$$L_1 \left( \frac{z_n - c_n}{d_n} \right) \rightarrow \mathcal{N}(0,1)$$

where

$$c_n = \int_{K_n^M} \ln g(t) \rho_1(dt), \quad d_n = \left( \int_{K_n^M} \ln^2 g(t) \rho_1(dt) \right)^{1/2},$$

$K_n^M$  is the same as in the Theorem 1.

From Theorem 1 and 2 we have

**THEOREM 3.** Assume that the following conditions are satisfied

1'')  $\rho_1 \approx \rho_0$

2)  $\int_{\Gamma} (f^{1/2}(t) - 1)^2 \rho_t(dt) < \infty$  where  $f(t) = \frac{d\rho_0}{d\rho_1}(t)$

3) There exists  $M > 0$  such that

$$(\rho_0 + \rho_1) \{x : |\ln f(x)| > M\} < \infty$$

Then as  $n \rightarrow \infty$

$$L_0 \left( \frac{z_n - a_n}{b_n} \right) \rightarrow \mathcal{N}(0,1), \quad L_1 \left( \frac{z_n - c_n}{d_n} \right) \rightarrow \mathcal{N}(0,1),$$

where  $a_n, b_n$  are the same as in Theorem 1 and  $c_n, d_n$  as in Theorem 2.

### III. THE CONVERGENCE RATE

**THEOREM 4.** *Assume that the conditions of the Theorem 1 are satisfied. Then there exists  $A > 0$  such that for every integer  $n$*

$$(11) \quad \sup_x \left| F_n(x) - \Phi(x) \right| \leq \frac{A}{b_n^{1/2}}$$

where

$$F_n(x) = Q_{\rho_0} \left\{ \mu : \frac{Z_n(\mu) - a_n}{b_n} < x \right\}$$

and

$$\Phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x \exp\left(-\frac{t^2}{2}\right) dt$$

**Proof:** Let us denote

$$F_n^{(1)}(x) = Q_{\rho_0} \left\{ \mu : \frac{Z_n^{(1)}(\mu) - a_n}{b_n} < x \right\}$$

Then for every  $\varepsilon > 0$  we have

$$(12) \quad F_n(x) - \Phi(x) = [F_n(x) - F_n^{(1)}(x + \varepsilon)] + [F_n^{(1)}(x + \varepsilon) - \Phi(x + \varepsilon)] + [\Phi(x + \varepsilon) - \Phi(x)].$$

Moreover, since

$$(13) \quad 0 < \Phi(x + \varepsilon) - \Phi(x) < \frac{1}{\sqrt{2\pi}} \varepsilon.$$

and

$$|F_n(x) - F_n^{(1)}(x + \varepsilon)| \leq Q_{\rho_0} \left\{ \left| \frac{Z_n^{(2)}}{b_n} \right| \geq \varepsilon \right\},$$

by Chebyshev's inequality we get

$$(14) \quad |F_n(x) - F_n^{(1)}(x + \varepsilon)| \leq \frac{A_1}{\varepsilon b_n}$$

where

$$A_1 = \rho_0 \{ |en_i| > M \}.$$

Now let  $\varphi_n(u)$  and  $\varphi(u)$  denote the characteristic functions of  $F_n^{(1)}$  and  $\Phi$ , respectively. Then we have

$$|\varphi_n(u) - \varphi(u)| = \varphi(u) \cdot \left| \exp\left(\psi_n(u) + \frac{u^2}{2}\right) - 1 \right|.$$

Hence and by (10)  $|u| \leq \left(\frac{b_n}{M}\right)^{1/3}$ , it follows that for any  $|u| < \left(\frac{b_n}{M}\right)^{1/3}$

$$|\varphi_n(u) - \varphi(u)| \leq 2\varphi(u) \left| \varphi_n(u) + \frac{u^2}{2} \right| \leq \frac{2M}{b_n} \cdot |u|^3 \cdot \exp \left\{ -\frac{u^2}{2} \right\}.$$

Therefore, by the Esseen's inequality (cf. [2], p. 207), we infer that for

$$\Gamma = \left( \frac{b_n}{M} \right)^{1/3}$$

$$(15) \sup_x |F_n^{(1)}(x) - \Phi(x)| \leq \frac{4}{\pi} \int_0^{\Gamma} \frac{M}{b_n} u^2 e^{-\frac{u^2}{2}} du + \frac{24}{\sqrt{2\pi^3} \cdot \Gamma} \leq \frac{A_2}{b_n} + \frac{A_3}{b_n^{1/3}},$$

where

$$A_2 = 4M/\pi, \quad A_3 = 24\sqrt[3]{M} / \sqrt{2\pi^3}.$$

From (12) + (15) the relation (11) holds with

$$A = \max(A_3, \sup_n (A_2/b_n^{2/3}), \sup_n (A_4/b_n^{1/6})$$

where

$$A_4 = \sqrt{A_1} / 2\sqrt{2\pi}.$$

Theorem 4 is proved.

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#### REFERENCES

- [1] K. Krickérberg, «Lectures on point processes» (in Vietnamese), Hanoi, 1976.
- [2] Yu. V. Prochorov, Yu. A. Rozanov, «Probability theory» (in Russian), Moscow, 1973.
- [3] E. L. Lehmann, «Testing statistical hypotheses» (in Russian), Moscow, 1964.